- **EX:** Are the bounds given by Chebyshev's inequality more accurate when f(x) is a uniform distribution or when f(x) is a gaussian distribution? Justify your answer.
  - **SOL'N:** Chebyshev's inequality is more helpful when a distribution has long tails. The probability density for a uniform distribution drops to zero for *x* more than a certain number of  $\sigma$ 's from the mean,  $\mu$ . For a uniform distribution from 0 to 1, for example,  $\sigma^2 = 1/12$ , and  $\sigma = 1/\sqrt{12}$ . The probability density drops to zero for values farther than 1/2 from  $\mu = 1/2$ . Solving  $c\sigma = 1/2$ , we find that  $c = \sqrt{3}$ . Thus, for a uniform distribution, we have

$$P(\mu - c\sigma \le X \le \mu + c\sigma) = 1 \text{ for } c \ge \sqrt{3}.$$

In this case, Chebyshev's inequality only guarantees a probability of

$$P(\mu - c\sigma \le X \le \mu + c\sigma) \ge 1 - \frac{1}{c^2} = \frac{2}{3}$$
 for  $c = \sqrt{3}$ .

Thus, Chebyshev's inequality is of little use for a uniform density function.

If we consider a standard gaussian (with  $\mu = 0$  and  $\sigma = 1$ ), the probability never drops to zero as we move away from the mean. If, for example, we consider  $c = \sqrt{3}$ , we can use a table for the area under a standard gaussian to find  $P(X \le \mu + c\sigma) = P(X \le \sqrt{3}) \approx P(X \le 1.73) = 0.9582$ . We subtract from this  $P(X \le \mu - c\sigma) = 0.0418$  to obtain

$$P(\mu - c\sigma \le X \le \mu + c\sigma) = 0.9582 - 0.0418 = 0.9164$$
 for  $c = \sqrt{3}$ .

In this case, Chebyshev's inequality guarantees a probability of

$$P(\mu - c\sigma \le X \le \mu + c\sigma) \ge 1 - \frac{1}{c^2} = \frac{2}{3} = 0.6667 \text{ for } c = \sqrt{3}.$$

This is better than the approximation for the uniform density function, although it still seems rather conservative.