THM: For any probability density function and any real number k > 0

$$P(|X-\mu| \ge k\sigma) \le \frac{1}{k^2} \text{ or } P(|X-\mu| \le k\sigma) \ge \frac{1}{k^2}$$

- **NOTE:** This theorem gives an upper bound on how much of the probability density can lie farther than $k\sigma$ from the mean value. Thus, the probability density is constrained in how far its tails can lie from the mean value on a scale measured by standard deviations.
- **NOTE:** This theorem is only useful for values of k > 1, since probability is always less than or equal to unity, and the theorem is most useful for larger values of k. For example, all but one-ninth of the probability lies within three standard deviations of the mean, regardless of what the probability density function happens to be.
- **PROOF:** We start with the definition of standard deviation:

$$\sigma^2 = \int_{-\infty}^{\infty} (x - \mu)^2 f(x) dx \, .$$

The figure below shows a generic probability density function, f(x).



For the calculation of σ^2 , we will multiply f(x) by the quadratic function $(x-\mu)^2$ added to the graph below.



The product $(x-\mu)^2 f(x)$ is shown below, and the area under this curve, (i.e., the integral of $(x-\mu)^2 f(x)$), shown in brown, is the value of σ^2 .



We split the integral for σ^2 into regions within $k\sigma$ of the mean (center region) and without $k\sigma$ of the mean (gray regions) giving us the following result.

$$\sigma^{2} = \int_{-\infty}^{\mu - k\sigma} (x - \mu)^{2} f(x) dx + \int_{\mu - k\sigma}^{\mu + k\sigma} (x - \mu)^{2} f(x) dx + \int_{\mu + k\sigma}^{\infty} (x - \mu)^{2} f(x) dx.$$

Since the quantities being integrated are all non-negative, if we were to delete the middle integral (i.e., the integral for values within $k\sigma$ of the mean) we would have the following result:

$$\sigma^2 \ge \int_{-\infty}^{\mu-k\sigma} (x-\mu)^2 f(x) dx + \int_{\mu+k\sigma}^{\infty} (x-\mu)^2 f(x) dx.$$

In other words, the areas under the side portions are less than the entire area. We obtain an even smaller area on the sides if we replace $(x - \mu)^2$ with a smaller multiplier, namely $k^2\sigma^2$. That is, for the integrals in the above equation we have $(x - \mu)^2 \ge k^2\sigma^2$, so we can write the following inequality:

$$\sigma^2 \ge \int_{-\infty}^{-k\sigma} k^2 \sigma^2 f(x) dx + \int_{k\sigma}^{\infty} k^2 \sigma^2 f(x) dx.$$

The figure below shows the right-hand side of this equation as red areas that are clearly smaller than the original side areas.



At this point, we factor out the $k^2\sigma^2$ from the integrals to obtain

$$\sigma^{2} \ge k^{2} \sigma^{2} \left(\int_{-\infty}^{-k\sigma} f(x) dx + \int_{k\sigma}^{\infty} f(x) dx \right)$$

or, if we divide both sides by σ ,

$$1 \ge k^2 \left(\int_{-\infty}^{-k\sigma} f(x) dx + \int_{k\sigma}^{\infty} f(x) dx \right).$$

The value in parentheses is now a probability, and we have

$$1 \ge k^2 P(|X - \mu| \ge k\sigma)$$

or, if we divide both sides by k^2 ,

$$\frac{1}{k^2} \ge P(|X-\mu| \ge k\sigma).$$

This result is equivalent to the theorem statement, and our proof is finished.

One might wonder the bound is achievable, and the answer for k > 1 is yes. The distribution shown below achieves the bound by putting as much of the probability as possible (i.e., $1/2k^2$) at points masses located at distance $k\sigma$ from μ . Thus, we have a discrete distribution:



We verify that the calculated variance is indeed σ^2 :

$$\sigma^{2} = \sum_{x_{i}} (x_{i} - \mu)^{2} P(x_{i}) = (-k\sigma)^{2} \frac{1}{2k^{2}} + 0 \cdot (1 - \frac{1}{k^{2}}) + (k\sigma)^{2} \frac{1}{2k^{2}},$$

which simplifies to $\sigma^2 = \sigma^2$, as required.

REF: Ronald E. Walpole, Raymond H. Myers, Sharon L. Myers, and Keying Ye, *Probability and Statistics for Engineers and Scientists*, 8th Ed., Upper Saddle River, NJ: Prentice Hall, 2007.